FROSTY: a high-dimensional scale-free Bayesian network learning method

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Introduction

Graphical models



Figure 1: Undirected Graph (UG), Directed Acyclic Graph (DAG), Completed Partially Directed Acyclic Graph (CPDAG)

Graphical models



Figure 1: Undirected Graph (UG), Directed Acyclic Graph (DAG), Completed Partially Directed Acyclic Graph (CPDAG)

$$UG: P(X) = \frac{1}{Z} \psi_1(X_1, X_2) \psi_2(X_2, X_3, X_4)$$

$$DAG: P(X) = \prod_{X_v \in \mathcal{V}} P(X_v \mid X_{pa(X_v)})$$

$$= P(X_1) P(X_2 \mid X_1) P(X_3 \mid X_2, X_4) P(X_4)$$

Bayesian networks

• Indistinguishablility (Markov equivalence class)

e.g., $V = \{X_1, X_2, X_3\}, I(P) = \{X_1 \perp X_3 \mid X_2\}$



Figure 2: Markov Equivalence Class

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Figure 2: Markov Equivalence Class

• Large non-convex search space (NP-hard)

 $\begin{aligned} |B_{p}| &= \sum_{i=1}^{p} (-1)^{i+1} {p \choose i} 2^{i(p-i)} |B_{p-i}| \\ |B_{14}| &= 1,439,428,141,044,398,334,941,790,719,839,535,103 \end{aligned}$

Recent works

- Annealing on Regularized Cholesky Score (ARCS)
 - developed by Ye, Amini, and Zhou (2020) from UCLA
 - $f(L; P) = -\log |L| + tr(PSP^TLL^T) + \sum_{i>j} \rho_{\theta}(L_{ij})$
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- Removal-Fill-Degree (RFD)
 - developed by Squires, Amaniampong, and Uhler (2020) from MIT
 - given undirected graph, recover the variable ordering based on removal/fill scores

Our Method: FROSTY

FROSTY

FROSTY utilizes Robust Selection (RobSel), (Cisneros, Petersen, and Oh, 2020) and sparse Choleksy algorithm called Approximate Minimum Degree ordering (AMD), (Amestoy, Davis, and Duff, 1996).

Algorithm 1 Pseudocode for FROSTY

Input : Dataset X, confidence level α

Output: Bayesian network B

Step 1: Estimate input undirected graph Θ

- 1: Set $\lambda = \mathsf{RobSel}(X, \alpha)$
- 2: Estimate Θ_{λ} with Graphical Lasso
- 3: Prune Θ_{λ} with conditional independence test

4: $\frac{Step \ 2}{L_{\pi}, P_{\pi}} = AMD(\Theta_{\lambda})$

5:
$$L = P_{\pi}^{T} L_{\pi} P_{\pi}$$
 and $B = (L_{D} - L) L_{D}^{-1}$

6: return B

Let $\ell(\Theta)$ be the Gaussian negative log-likelihood. Cisneros, Petersen, and Oh (2020) show that

$$\underbrace{\min_{\Theta} \sup_{\mathcal{P}: D_{c}(\mathcal{P}, \mathcal{P}_{n}) \leq \lambda}}_{\text{Distributionally Robust}} E_{\mathcal{P}}[\ell(\Theta)] = \underbrace{\min_{\Theta} \{\ell(\Theta) + \lambda \|\Theta\|_{1}\}}_{\text{Graphical Lasso formulation}},$$
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Under the DRO formulation, the *optimal* regularization parameter can be chosen by

$$\lambda = \inf \{\lambda > 0 : P_0(\Theta \in \mathcal{C}_n(\lambda)) \ge 1 - \alpha\}.$$
 (2)

This can be efficiently computed by bootstrapping dataset X only.

The interpretation of α in RobSel is recently shown in Tran et al. (2022): $P(\text{having at least one false positive edge in } \hat{\Theta}) \leq \alpha \text{ as } n \to \infty$ (3) The interpretation of α in RobSel is recently shown in Tran et al. (2022): $P(\text{having at least one false positive edge in } \hat{\Theta}) \leq \alpha \text{ as } n \to \infty$ (3)

In other words, the tuning parameter α in RobSel has a direct connection to the asymptotic family-wise error rate of the zero-nonzero patterns.

Under multivariate Gaussian assumption, $\Theta = \Sigma^{-1}$ corresponds to the undirected graph, which can be expressed as a function of *B* and $\Omega = diag(\omega_1^2, \ldots, \omega_p^2)$, the error variances of X_1, \ldots, X_p :

$$\Theta(B,\Omega) = (I-B)\Omega^{-1}(I-B)^T$$
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If we appropriately order the variables, say $\pi = (\pi_{(1)}, \ldots, \pi_{(p)})$,

$$\begin{aligned} \Theta_{\pi}(B_{\pi},\Omega_{\pi}) &= (I - B_{\pi})\Omega_{\pi}^{-1}(I - B_{\pi})^{T} \\ &= LDL^{T} \\ &= \widetilde{L}\widetilde{L}^{T} \end{aligned}$$
(5)

Approximate minimum degree ordering



Figure 3: The impact of variable ordering.

- RobSel: $O(bnp^2)$
- Graphical lasso: $O(p^3)$
- CI-testing: $O(ns^2 + s^3)$
- AMD: approximately $O(|L|) \ll O(p^2)$

where *b* is the number of boostrap samples, *s* is the average number of nonzeros per row in Θ , and *L* is the Cholesky factor.

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Note: one only needs to fit Graphical Lasso once.

Simulation: empirical results

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• Matthew's Correlation Coefficient, $-1 \leq \mathsf{MCC} \leq 1$

$$MCC = \frac{TP \times TN - FP \times FN}{\sqrt{(TP + FP)(TP + FN)(TN + FP)(TN + FN)}}$$

Simulation results



Figure 4: Performance comparison for scale-free graphs.

Simulation results



Figure 5: Performance comparison for Erdos-Renyi graphs.

Simulation results

Figure 6: Performance comparison for Pathfinder network.

Method	Runtime (sec.)				
	p=100	p=200	p=500	p=1000	p=2000
FROSTY	0.12	0.28	1.39	5.52	33.69
RFD (RS)	0.86	6.52	168.62	2180.04	-
RFD (CV)	33.59	177.88	1817.85	14250.32	-
ARCS	57.52	101.47	562.56	5721.50	14581.00

Note: For p = 2000, RFD took up too much memory that it was not feasible to run within our available memory space of 60GB.

Table 1: Runtime analysis

Conclusion

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- FROSTY is simple.
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- First step of FROSTY can improve other methods.
 - Methods that take an undirected graph as an input can be significantly improved by adapting FROSTY's undirected graph estimation step.

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Appendix

$$\begin{split} \lambda &= \inf \left\{ \lambda > 0 : P_0(\Theta \in \mathcal{C}_n(\lambda)) \ge 1 - \alpha \right\} \\ &= \inf \left\{ \lambda > 0 : P_0(R_n(\Theta) \le \lambda) \ge 1 - \alpha \right\} \\ &= \inf \left\{ \lambda > 0 : P_0(\|\operatorname{vec}(S - \Theta^{-1})\|_{\infty} \le \lambda) \ge 1 - \alpha \right\} \end{split}$$
(6)

where $C_n(\lambda)$ is the confidence region for Θ and

$$R_{n}(\Theta) = \inf \left\{ D_{c}(\mathcal{P}, \mathcal{P}_{n}) : \\ E_{\mathcal{P}}\left[\frac{\partial}{\partial \Theta'} (tr(S\Theta') - \log |\Theta'|) \Big|_{\Theta' = \Theta} \right] = \mathbf{0} \right\}$$
(7)

is called the Robust Wasserstein Profile (RWP) function, which represents the minimum distance between the empirical distribution and any *plausible* distribution that satisfies the first order optimality condition for the precision matrix Θ .

Undirected graph estimation

Figure 7: Undirected graph estimation comparison.