Convergence Complexity of Gibbs Samplers for Bayesian Vector Autogressions

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Bayesian VARX

r-valued stochastic process

$$Y_t = \sum_{i=1}^q \mathcal{A}_i^T Y_{t-i} + \mathcal{B}^T X_t + \varepsilon_t \qquad \varepsilon_t \stackrel{ind}{\sim} N_r(0, \Sigma)$$

 $\{X_t\}$ indep. $\{\varepsilon_t\}$ and distribution not depending on $\{A_i\}, \mathcal{B}, \Sigma$

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Rewrite

$$Y_t = \mathcal{A}^T Z_t + \mathcal{B}^T X_t + \varepsilon_t$$

$$\mathcal{A} = [\mathcal{A}_1^T, \dots, \mathcal{A}_q^T] \in \mathbb{R}^{qr imes r}, \quad \mathcal{B} \in \mathbb{R}^{p imes r}$$

 $Z_t = [Y_{t-1}^T, \dots, Y_{t-q}^T]^T \in \mathbb{R}^{qr}, \quad Z_1 \text{ is fixed}$

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Rewrite

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$$egin{aligned} \mathcal{A} &= [\mathcal{A}_1^{\mathcal{T}}, \dots, \mathcal{A}_q^{\mathcal{T}}] \in \mathbb{R}^{qr imes r}, \quad \mathcal{B} \in \mathbb{R}^{p imes r} \ \mathcal{Z}_t &= [Y_{t-1}^{\mathcal{T}}, \dots, Y_{t-q}^{\mathcal{T}}]^{\mathcal{T}} \in \mathbb{R}^{qr}, \quad \mathcal{Z}_1 ext{ is fixed} \end{aligned}$$

small-*n*: Data fixed, n > p but possibly small compared to qrlarge-*n*: Data stochastic, n > p, and *n* is increasing

Likelihood

$$Y_t = \mathcal{A}^T Z_t + \mathcal{B}^T X_t + \varepsilon_t \qquad \varepsilon_t \stackrel{ind}{\sim} N_r(0, \Sigma)$$

$$S = n^{-1}(Y_t - \mathcal{A}^T Z_t - \mathcal{B}^T X_t)^T (Y_t - \mathcal{A}^T Z_t - \mathcal{B}^T X_t)$$

$$f(Y, X | \mathcal{A}, \mathcal{B}, \Sigma) \propto |\Sigma|^{-n/2} \operatorname{etr} \left(-\frac{n}{2} S \Sigma^{-1} \right)$$

Instead of \mathcal{A} it is common to work with $\alpha = \text{vec}(\mathcal{A}) \in \mathbb{R}^{qr^2}$. Thus we need a prior for

$$(\alpha, \mathcal{B}, \Sigma) \in \mathbb{R}^{qr^2} imes \mathbb{R}^{p imes r} imes S^r_{++}$$

VARX Priors

$$Y_t = \mathcal{A}^T Z_t + \mathcal{B}^T X_t + \varepsilon_t \qquad \varepsilon_t \stackrel{ind}{\sim} N_r(0, \Sigma)$$

Karlsson (2013, Handbook Economic Forecasting) gives a comprehensive review of priors.

Let
$$\Psi = [\mathcal{A}^T, \mathcal{B}^T]^T$$
.
vec $(\Psi) \sim MVN$ and $\Sigma \sim$ Inverse Wishart,
 $f(\Psi, \Sigma) \propto |\Sigma|^{-a}$,

Minnesota prior, and so on.

Proposed prior: Recall $\alpha = \operatorname{vec}(\mathcal{A}) \in \mathbb{R}^{qr^2}$

$$f(\alpha)$$
 $f(\mathcal{B}) \propto 1$ $f(\Sigma) \propto |\Sigma|^{-a/2} \operatorname{etr} \left(-\frac{1}{2}D\Sigma^{-1}\right)$

VARX Priors

$$\alpha = \mathsf{vec}(\mathcal{A}) \in \mathbb{R}^{qr^2}$$
 and

$$f(lpha) \qquad f(\mathcal{B}) \propto 1 \qquad f(\Sigma) \propto |\Sigma|^{-a/2} \mathsf{etr}\left(-\frac{1}{2}D\Sigma^{-1}\right)$$

<u>Thm</u> If either

- 1. *D* is positive definite, *X* has full column rank, n + a > 2r + p, and $f(\alpha)$ is proper; or
- 2. [Y, Z, X] has full column rank, n + a > (2 + q)r + p, and $f(\alpha)$ is bounded,

then the posterior $f(\alpha, \mathcal{B}, \Sigma | \mathcal{D}_n)$ exists and is proper.

VARX Posterior

$$\begin{split} f(\mathcal{B}) \propto 1 & f(\Sigma) \propto |\Sigma|^{-a/2} \mathrm{etr}\left(-\frac{1}{2}D\Sigma^{-1}\right) \\ f(\alpha) \propto \exp\left(-\frac{1}{2}(\alpha-m)^{T}C(\alpha-m)\right) \end{split}$$

 $f(\alpha)$ common in macroeconomics and finance allows large VARs, i.e. qr large and m can be chosen to address near non-stationarity (unit root sense) $f(\mathcal{B})$ common in multivariate location-scale settings

 $f(\Sigma)$ includes inverse Wishart and Jeffreys priors

Collapsed Gibbs sampler

 (α, Σ) is a linchpin variable:

$$f(\alpha, \mathcal{B}, \Sigma | \mathcal{D}_n) = f(\mathcal{B} | \alpha, \Sigma, \mathcal{D}_n) f(\alpha, \Sigma | \mathcal{D}_n)$$

and

$$\mathcal{B} \mid \alpha, \Sigma, \mathcal{D}_n \sim \mathsf{Matrix} \mathsf{Normal}$$

Use Gibbs sampler for $f(\alpha, \Sigma | D_n)$ since

 $\Sigma \mid \alpha, \mathcal{D}_n \sim \text{Inverse Wishart}$ $\alpha \mid \Sigma, \mathcal{D}_n \sim \text{Multivariate Normal}$

$$\theta = (\mathcal{B}, \Sigma, \alpha) \to (\mathcal{B}, \Sigma', \alpha) \to (\mathcal{B}, \Sigma', \alpha') \to (\mathcal{B}', \Sigma', \alpha') = \theta'$$

Convergence Analysis

Geometric ergodicity of Collapsed Gibbs sampler: Find $\rho < 1 \mbox{ s.t.}$

$$\|\mathcal{K}_{\mathcal{C}}^{h}(\theta,\cdot)-\mathcal{F}(\cdot|\mathcal{D}_{n})\|_{\mathcal{T}V}\leq M(\theta)\rho^{h}$$

Classical (small *n*):

Find conditions to ensure that geometric ergodicity holds for any fixed data set.

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Classical (small *n*):

Find conditions to ensure that geometric ergodicity holds for any fixed data set.

Convergence complexity (large *n*): $\rho = \rho_n$ Find conditions to ensure that the convergence rate behaves well for large *n*:

$$\limsup_{n \to \infty} \rho_n < 1 \quad \text{ almost surely}$$

so the geometric ergodicity is *asymptotically stable*.

Collapsed Gibbs sampler:

$$\theta = (\mathcal{B}, \Sigma, \alpha) \to (\mathcal{B}, \Sigma', \alpha) \to (\mathcal{B}, \Sigma', \alpha') \to (\mathcal{B}', \Sigma', \alpha') = \theta'$$

It suffices to study the marginal process $\{\alpha^h\}$ because

$$\|K_{\mathcal{C}}^{h}(\theta,\cdot) - F(\cdot|\mathcal{D}_{n})\|_{\mathcal{T}V} \leq \|K_{\alpha}^{h-1}(\alpha,\cdot) - F_{\alpha}(\cdot|\mathcal{D}_{n})\|_{\mathcal{T}V}$$

Convergence Analysis

Rosenthal (JASA, 1995)

Suppose $V: \mathbb{R}^{qr^2}
ightarrow [0,\infty), \ \lambda < 1$ and some $L < \infty$

$$\int V(\alpha) K_{\alpha}(\alpha', d\alpha) \leq \lambda V(\alpha') + L \quad \text{ for all } \alpha'.$$
 (1)

and for
$$T > 2L/(1 - \lambda)$$

 $K_{\alpha}(\alpha, \cdot) \ge \varepsilon R(\cdot)$ for all $\alpha \in \{\alpha : V(\alpha) \le T\}.$ (2)

Then K_{α} is geometrically ergodic and there is an explicit formula for

$$\bar{\rho} = \bar{\rho}_n(Y, X, C, D, m, a)$$

such that

$$\rho \leq \bar{\rho}$$

Convergence Analysis

Take home message:

Rosenthal's theorem yields an explicit upper bound on the rate

$$\bar{\rho} = \bar{\rho}(\mathcal{D}_n, C, D, m, a),$$

but if V is not chosen carefully, then it is likely the case that

 $\liminf_{n\to\infty}\bar\rho\to 1 \quad \text{ almost surely} \quad$

and we won't be able to conclude that the geometric ergodicity is asymptotically stable.

Recall

$$f(\alpha) \propto \exp\left(-\frac{1}{2}(\alpha-m)^{T}C(\alpha-m)\right)$$

<u>Thm</u> If C is positive definite, then the collapsed Gibbs sampler is geometrically ergodic.

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Used drift function

$$V(\alpha) = \|\alpha\|^2$$

which implies that the Markov chain should visit sets near the origin frequently.

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No reason to think this is reasonable when n is large.

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Maybe we should center the drift function around the least squares estimator or MLE.

$$V(\alpha) = \|Q_X Z A - Q_X Z \hat{A}\|_F^2$$
$$= \|(I_r \otimes Q_X Z)(\alpha - \hat{\alpha})\|^2$$

Convergence Complexity (large *n*) Analysis <u>Thm</u> If

- (a) C is positive definite,
- (b) [Y, X, Z] has full column rank for all large enough n almost surely,

(c)
$$\|\hat{lpha}\|^2 = O(1)$$
 almost surely as $n o \infty$, and

(d) there exists $0 \le M < \infty$ such that, almost surely,

$$M^{-1} \leq \liminf_{n \to \infty} n^{-1} \lambda_{\min}(Y^T Q_{[Z,X]} Y)$$

$$\leq \limsup_{n \to \infty} n^{-1} \lambda_{\max}(Y^T Q_{[Z,X]} Y)$$

$$\leq M$$

then, almost surely,

 $\limsup_{n\to\infty}\bar{\rho}_n<1$