## An R Package for Linear Mediation Analysis with Complex Survey Data

Yujiao Mai, Jiahui Xu, Deo Kumar Srivastava, Hui Zhang Department of Biostatistics

## Outline



1. Classic Mediation Analysis
2. Complex Surveys
3. Problem \& Solution
4. Software \& Application

## Section 1

## Classic Mediation Analysis



## Questions to answer

Regression analysis:
"Does it work?"

Mediation analysis: "How does it work?"



## Model

(Continuous Variables $M, Y$ )


- Estimator: $F_{M L}\left(\theta ; S_{n}\right)=\log |\Sigma(\theta)|+\operatorname{trace}\left[S_{n} \Sigma(\theta)^{-1}\right]-\log \left|S_{n}\right|-p$
$S_{n}$ : sample covariance matrix
$\Sigma(\theta)$ : estimated covariance matrix, $\theta=\left(\mu_{M}, \mu_{Y}, \alpha, \gamma, \beta, \sigma_{\varepsilon_{M}}^{2}, \sigma_{\varepsilon_{Y}}^{2}\right), p$ : number of parameters

$$
\operatorname{VAR}(M ; \theta)=\operatorname{VAR}\left(\mu_{M}+\alpha X\right)+\sigma_{\varepsilon_{M}}^{2}
$$

## Model (Continue)

- Significance Test: $\mathrm{H}_{0}: \alpha \beta=0 ; \mathrm{H}_{1}: \alpha \beta \neq 0$

1) Sobel's Test (Sobel, 1982)

$$
T_{\text {Sobel }}=\hat{\alpha} \hat{\beta} / \sqrt{\hat{\beta}^{2} \sqrt{\operatorname{Gar}}(\hat{\alpha})+\hat{\alpha}^{2} \sqrt{\operatorname{Gar}}(\hat{\beta})} \quad H_{0}^{\sim} N(0,1) \text { as } n \rightarrow \infty
$$

2) Resampling/bootstrap

## Assumptions

True Model
Independent Residuals
Normality
Independent-identically-distributed sample data

## Section 2

Complex Surveys

## Multi-stage <br> Sampling (Wolerer 2007)

Suppose the population ( 40,000 students) locate in two states.

Each state has 20 districts.
Each district has 10 schools.
Each school has 100 students.

Sample weights $\left(\omega_{0}\right)=$ the number of students represented

| State | District | School | Student | $\mathbf{Y}$ | $\boldsymbol{\omega}_{0}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 1 | 1 | 9.8 | 2500 |
| 1 | 1 | 1 | 2 | 7.5 | 2500 |
| 1 | 1 | 2 | 3 | 8.3 | 2500 |
| 1 | 1 | 2 | 4 | 4.5 | 2500 |
| 1 | 2 | 3 | 5 | 4.5 | 2500 |
| 1 | 2 | 3 | 6 | 5.1 | 2500 |
| 1 | 2 | 4 | 7 | 2.3 | 2500 |
| 1 | 2 | 4 | 8 | 6.5 | 2500 |
| 2 | 3 | 5 | 9 | 8.1 | 2500 |
| 2 | 3 | 5 | 10 | 3.2 | 2500 |
| 2 | 3 | 6 | 11 | 4.5 | 2500 |
| 2 | 3 | 6 | 12 | 5.8 | 2500 |
| 2 | 4 | 7 | 13 | 6.6 | 2500 |
| 2 | 4 | 7 | 14 | 8.1 | 2500 |
| 2 | 4 | 8 | 15 | 1.2 | 2500 |
| 2 | 4 | 8 | 16 | 6.5 | 2500 |

$$
\omega_{0}=\frac{1}{\pi_{\mathrm{PSU}} \times \pi_{\text {sch } \mid \mathrm{PSU}} \times \pi_{\text {stud } \mid \text { sch }}}=\frac{1}{\frac{2}{20} \times \frac{2}{10} \times \frac{2}{100}}=2,500
$$

## Problem: Estimate the Mean of $Y$

- Disaggregated estimates vs. aggregated estimates
(Group-specific effects)
(Generalized effects)
For aggregated estimates:
- Point estimate is consistent when including the sample weights
- Standard errors is underestimated even when taking into account the sample weights


## Adjustments:

Taylor series linearization (TSL)
Bootstrap
Jackknife repeated replications (JRR)
Balanced repeated replications (BRR)

## Balanced repeated replications

| State | District | School | Student | Y | $\omega_{0}$ | $\omega_{1}^{\prime}$ | $\omega_{2}^{\prime}$ | $\boldsymbol{\omega}_{3}^{\prime}$ | $\boldsymbol{\omega}_{4}^{\prime}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 1 | 1 | 9.8 | 2500 | 0 | $2 \times 2500$ | 0 | $2 \times 2500$ |
| 1 | 1 | 1 | 2 | 7.5 | 2500 | 0 | $2 \times 2500$ | 0 | $2 \times 2500$ |
| 1 | 1 | 2 | 3 | 8.3 | 2500 | 0 | $2 \times 2500$ | 0 | $2 \times 2500$ |
| 1 | 1 | 2 | 4 | 4.5 | 2500 | 0 | $2 \times 2500$ | 0 | $2 \times 2500$ |
| 1 | 2 | 3 | 5 | 4.5 | 2500 | $2 \times 2500$ | 0 | $2 \times 2500$ | 0 |
| 1 | 2 | 3 | 6 | 5.1 | 2500 | $2 \times 2500$ | 0 | $2 \times 2500$ | 0 |
| 1 | 2 | 4 | 7 | 2.3 | 2500 | $2 \times 2500$ | 0 | 2×2500 | 0 |
| 1 | 2 | 4 | 8 | 6.5 | 2500 | $2 \times 2500$ | 0 | $2 \times 2500$ | 0 |
| 2 | 3 | 5 | 9 | 8.1 | 2500 | 0 | 0 | 2×2500 | $2 \times 2500$ |
| 2 | 3 | 5 | 10 | 3.2 | 2500 | 0 | 0 | 2×2500 | $2 \times 2500$ |
| 2 | 3 | 6 | 11 | 4.5 | 2500 | 0 | 0 | 2×2500 | $2 \times 2500$ |
| 2 | 3 | 6 | 12 | 5.8 | 2500 | 0 | 0 | 2×2500 | 2×2500 |
| 2 | 4 | 7 | 13 | 6.6 | 2500 | $2 \times 2500$ | $2 \times 2500$ | 0 | 0 |
| 2 | 4 | 7 | 14 | 8.1 | 2500 | $2 \times 2500$ | $2 \times 2500$ | 0 | 0 |
| 2 | 4 | 8 | 15 | 1.2 | 2500 | $2 \times 2500$ | $2 \times 2500$ | 0 | 0 |
| 2 | 4 | 8 | 16 | 6.5 | 2500 | $2 \times 2500$ | $2 \times 2500$ | 0 | 0 |


Replicate sampling
weights $\boldsymbol{\omega}^{\prime}{ }_{r}, r=1,2, \ldots, R$
$R$ is the number of replications, $R=4$ in the case.
$\hat{\mu}_{r}$ is the estimate using replicate weight $\boldsymbol{w}_{r}^{\prime}$.
$\hat{\mu}$ is the estimate using original(main) sample weights

$$
S E_{B R R}(\hat{\mu})=\sqrt{\frac{\sum_{r=1}^{R}\left(\hat{\mu}_{r}-\hat{\mu}\right)^{2}}{R}}
$$

Section 3
Problem \& Solution


## Problem to Solve

## How can mediation analysis work with complex surveys?

| Classic Mediation Analysis | Complex Surveys |
| :---: | :---: |
| i-i-d sample: |  |
| Independent | Within cluster: <br> Dependent <br> i-i-d sample: <br> Identically distributed |
| Between cluster: <br> Heterogenous |  |
| U-i-d sample: <br> Equal possibility | Disaggregated <br> Estimates <br> Unequal-sized strata: <br> Unequal possibility |

## Potential Solutions

- Aggregated Estimates

Design-based method:

- Taylor series linearization (TSL)
- Bootstrap
- Jackknife repeated replications (JRR)
- Balanced repeated replications (BRR)

Most national/international surveys do not provide the cluster indicator.

## Apply Balanced Repeated Replications to Analysis of $\alpha \beta$

## Estimator

Maximum Likelihood $F_{M L}\left(\theta ; S_{n}\right)=\log |\Sigma(\theta)|+\operatorname{trace}\left[S_{n} \Sigma(\theta)^{-1}\right]-\log \left|S_{n}\right|-p$

Weighted covariance matrix

Replace $S_{n}$ with weighted sample covariance matrix $S_{w n}$

## Test Statistic

Asymptotical Normality (Bishop, 1975; Rao, 1973) $\hat{\alpha} \hat{\beta} \sim N\left(\mu, \sigma^{2}\right)$ as $\mathrm{n} \rightarrow \infty$

Use BRR Standard Errors

$$
T_{\mathrm{BRR}}=\frac{\hat{\alpha} \hat{\beta}}{\mathrm{SE}_{\mathrm{BRR}}(\hat{\alpha} \hat{\beta})} \stackrel{H_{0}}{\sim} N(0,1) \text { as } \mathrm{n} \rightarrow \infty
$$

$$
\mathrm{SE}_{\mathrm{BRR}}(\hat{\alpha} \hat{\beta})=\sqrt{\frac{1}{R(1-f)^{2}} \sum_{r=1}^{R}\left(\hat{\alpha}_{r} \hat{\beta}_{r}-\hat{\alpha}_{0} \hat{\beta}_{0}\right)^{2}}
$$

## Section 4

Software \& Application

## Software Packages

- R package 'MedSurvey’ (Feb. 2019)

Flexible and complex models
https://CRAN.R-project.org/package=MedSurvey

## Application

## Data:

2014-15 CPS Tobacco Use Supplement (TUS;
U.S. Department of Commerce and U.S. Census Bureau 2016), employed adult daily smokers (Non-Hispanic White males only).

Survey Design:
Balanced Repeated


Replications $R=160$.
$f=0.5$ is suggested.

```
## Package and options
1ibrary("MedSurvey")
## Data and related information
MedData
R <- 160
wgtnames <- paste("repwgt", seq(0,R,by=1), sep="")
mwgtname=wgtnames[1]
repwgtnames=wgtnames[2:(R+1)]
```

```
## Sepcify the model 1
mode11 <- '
    numcg ~ u0*1 + gamma0*workban + b1*sp_ad7tban + b2*sp_kidsban
    sp_ad7tban ~ u1*1 + a1*workban
    sp_kidsban ~ u2*1 + a2*workban
    sp_adltban ~~ sp_kidsban
    a1b1 := a1*b1
    a2b2 := a2*b2
    tota1 := gamma0 + (a1*b1) + (a2*b2)
```

\#\# Fit the mode1 1
fit. BRR <- med.fit. BRR(mode1=mode11, data=MedData, mwgtname=mwgtname,
repwgtnames=repwgtnames, fayfactor=0.5, paral1e1 = 'para11e1')
|\#\# view the summary results of the mediation analysis
med.summary (fit=fit.BRR, med.eff=c('a1b1', 'a2b2'))

## Results

Table 1. Tests for Mediation Effects between Smoking Ban at Work and Number of Cigarettes Smoked per Day among Male NH White Daily Smokers

| Mediator | Not Weighted |  |  |  | BRR Weighted |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\hat{\alpha} \widehat{\beta}$ | $\widehat{S D}$ | $p$-value | Adjusted $p$-value | $\hat{\alpha} \hat{\beta}$ | $\widehat{S D}_{\text {BRR }}$ | $p$-value | Adjusted $p$-value |
| $M_{1}$ | -0.0170 | 0.0066 | 0.010 | 0.020 | -0.0156 | 0.0073 | 0.032 | 0.063 |
| $M_{2}$ | 0.0018 | 0.0033 | 0.588 | 0.612 | 0.0004 | 0.0045 | 0.937 | 1.000 |
| ote. $\begin{aligned} & M 1=\text { Sup } \\ & M 2=\text { Sup } \\ & n=2,260 \\ & \text { Adjusted } \end{aligned}$ | orting Smok rting Smokin stimated po values adjus | ng Ban in ing Ban in ulation s ed for th | Adult-excl Children's ize is 3,29 multiple |  | Holm's me | hod (Holm | ^,1979). |  |

## Software Packages

- R package 'MedSurvey’ (Feb. 2019)

Flexible and complex models
https://CRAN.R-project.org/package=MedSurvey

- R shiny (March 2019)

User-friendly interface
https://sjbiostat.shinyapps.io/MedSurvey/

## R shiny




## Summary Download

Multimediation with Complex Survey Data

|  | Effect | Estimate | BRR SE. | P-Value | Adjusted P-Value |
| :--- | :--- | :--- | :--- | :--- | :--- |
| sp_adltban | a1b1 | -0.015456 | 0.004574 | 0.000728 | 0.001456 |
| sp_kidsban | a2b2 | -0.005366 | 0.003852 | 0.1636 | 0.1636 |

$p$ Value adjustment method is holm
Standard errors type is BRR SE


Look, this's our product.

## Software Packages

- R package 'MedSurvey’ (Feb. 2019)

Flexible and complex models
https://CRAN.R-project.org/package=MedSurvey

- R shiny (March 2019)

User-friendly interface
https://sibiostat.shinyapps.io/MedSurvey/

- SAS macros 'MedBRR' (Dec. 2018)

Super large-scale datasets
https://github.com/YuiiaoMai/MedSurvey

## References

- Bishop, Y. M., Fienberg, S. E., \& Holland, P. W. (1975). Discrete multivariate analysis: theory and practice. Cambridge, Massachusetts: Cambridge, Mass., MIT Press.
- Fay, R. E., \& Train, G. F. (1995). Aspects of survey and model-based postcensal estimation of income and poverty characteristics for states and counties. In Proceedings of the Section on Government Statistics, American Statistical Association, Alexandria, VA (pp. 154-159).
- Judkins, D. R. (1990). Fay's method for variance estimation. Journal of Official Statistics, 6(3), 223-239.
- Mai, Y., Ha, T., \& Soulakova, J. N. (2019). Multimediation Method With Balanced Repeated Replications For Analysis Of Complex Surveys. Structural Equation Modeling: A Multidisciplinary Journal.
- Rao, C. R. (1973). Linear statistical inference and its applications (2nd ed., Vol. 2). New York, NY: John Wiley \& Sons, Inc.
- Sobel, M. E. (1982). Asymptotic confidence intervals for indirect effects in structural equation models. Sociological Methodology, 13, 290-312.
- U.S. Department of Commerce, \& U.S. Census Bureau. (2016). National Cancer Institute and Food and Drug Administration co-sponsored Tobacco Use Supplement to the Current Population Survey. 201415.
- Wolter, K. (2007). Introduction to variance estimation. New York, NY: Springer.


## Acknowledgements

- The research is sponsored by American Lebanese Syrian Associated Charities (ALSAC).


St. Jude Childrens Research Hospital

Finding cures. Saving children


Thank you!

